# An Introduction to the Wavelet Analysis of Time Series

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### Overview

- wavelets are analysis tools mainly for
  - time series analysis (focus of this tutorial)
  - image analysis (will not cover)
- as a subject, wavelets are
  - relatively new (1983 to present)
  - synthesis of many new/old ideas
  - keyword in 10, 558+ articles & books since 1989
    (2000+ in the last year alone)
- broadly speaking, have been two waves of wavelets
  - continuous wavelet transform (1983 and on)
  - discrete wavelet transform (1988 and on)

# Game Plan

- introduce subject via CWT
- describe DWT and its main 'products'
  - multiresolution analysis (additive decomposition)
  - analysis of variance ('power' decomposition)
- $\bullet$  describe selected uses for DWT
  - wavelet variance (related to Allan variance)
  - decorrelation of fractionally differenced processes (closely related to power law processes)
  - signal extraction (denoising)

### What is a Wavelet?

- wavelet is a 'small wave' (sinusoids are 'big waves')
- real-valued  $\psi(t)$  is a wavelet if
  - 1. integral of  $\psi(t)$  is zero:  $\int_{-\infty}^{\infty} \psi(t) dt = 0$
  - 2. integral of  $\psi^2(t)$  is unity:  $\int_{-\infty}^{\infty} \psi^2(t) dt = 1$  (called 'unit energy' property)
- wavelets so defined deserve their name because

- #2 says we have, for every small  $\epsilon > 0$ ,

$$\int_{-T}^{T} \psi^2(t) \, dt < 1 - \epsilon,$$

for some finite T (might be quite large!)

- length of [-T, T] small compare to  $[-\infty, \infty]$
- #2 says  $\psi(t)$  must be nonzero somewhere
- -#1 says  $\psi(t)$  balances itself above/below 0
- Fig. 1: three wavelets
- Fig. 2: examples of complex-valued wavelets

### **Basics of Wavelet Analysis: I**

- wavelets tell us about variations in local averages
- to quantify this description, let x(t) be a 'signal'
  - real-valued function of t
  - will refer to t as time (but can be, e.g., depth)
- consider average value of x(t) over [a, b]:

$$\frac{1}{b-a} \int_{a}^{b} x(u) \, du \equiv \alpha(a,b)$$

• reparameterize in terms of  $\lambda \& t$ 

$$A(\lambda,t) \equiv \alpha(t-\frac{\lambda}{2},t+\frac{\lambda}{2}) = \frac{1}{\lambda} \int_{t-\frac{\lambda}{2}}^{t+\frac{\lambda}{2}} x(u) \, du$$

- $-\lambda \equiv b-a$  is called scale
- -t = (a+b)/2 is center time of interval
- $A(\lambda, t)$  is average value of x(t) over scale  $\lambda$  at t

#### **Basics of Wavelet Analysis: II**

- average values of signals are of wide-spread interest
  - hourly rainfall rates
  - monthly mean sea surface temperatures
  - yearly average temperatures over central England
  - etc., etc., etc. (Rogers & Hammerstein, 1951)
- Fig. 3: fractional frequency deviates in clock 571
  - can regard as averages of form  $\left[t \frac{1}{2}, t + \frac{1}{2}\right]$
  - -t is measured in days (one measurment per day)
  - plot shows A(1,t) versus integer t
  - $-A(1,t) = 0 \Rightarrow$  master clock & 571 agree perfectly
  - $-A(1,t) < 0 \Rightarrow$  clock 571 is losing time
  - can easily correct if A(1,t) constant
  - quality of clock related to changes in A(1,t)

# **Basics of Wavelet Analysis: III**

• can quantify changes in A(1,t) via

$$\begin{split} D(1,t-\frac{1}{2}) &\equiv A(1,t) - A(1,t-1) \\ &= \int_{t-\frac{1}{2}}^{t+\frac{1}{2}} x(u) \, du - \int_{t-\frac{3}{2}}^{t-\frac{1}{2}} x(u) \, du, \end{split}$$

or, equivalently,

$$D(1,t) = A(1,t+\frac{1}{2}) - A(1,t-\frac{1}{2})$$
  
=  $\int_{t}^{t+1} x(u) \, du - \int_{t-1}^{t} x(u) \, du$ 

• generalizing to scales other than unity yields

$$\begin{array}{lll} D(\lambda,t) &\equiv & A(\lambda,t+\frac{\lambda}{2}) - A(\lambda,t-\frac{\lambda}{2}) \\ &= & \frac{1}{\lambda} \int_{t}^{t+\lambda} x(u) \, du - \frac{1}{\lambda} \int_{t-\lambda}^{t} x(u) \, du \end{array}$$

- $D(\lambda, t)$  often of more interest than  $A(\lambda, t)$
- can connect to Haar wavelet: write

$$D(\lambda, t) = \int_{-\infty}^{\infty} \tilde{\psi}_{\lambda, t}(u) x(u) \, du$$

with

$$\tilde{\psi}_{\lambda,t}(u) \equiv \begin{cases} -1/\lambda, & t - \lambda \leq u < t; \\ 1/\lambda, & t \leq u < t + \lambda; \\ 0, & \text{otherwise.} \end{cases}$$

## **Basics of Wavelet Analysis: IV**

• specialize to case 
$$\lambda = 1$$
 and  $t = 0$ :

$$\tilde{\psi}_{1,0}(u) \equiv \begin{cases} -1, & -1 \le u < 0; \\ 1, & 0 \le u < 1; \\ 0, & \text{otherwise.} \end{cases}$$

comparison to  $\psi^{\mathrm{H}}(u)$  yields  $\tilde{\psi}_{1,0}(u) = \sqrt{2}\psi^{\mathrm{H}}(u)$ 

• Haar wavelet mines out info on difference between unit scale averages at t = 0 via

$$\int_{-\infty}^{\infty} \psi^{\mathrm{H}}(u) x(u) \, du \equiv W^{\mathrm{H}}(1,0)$$

• to mine out info at other t's, just shift  $\psi^{H}(u)$ :

$$\psi_{1,t}^{\mathrm{H}}(u) \equiv \psi^{\mathrm{H}}(u-t); \text{ i.e., } \psi_{1,t}^{\mathrm{H}}(u) = \begin{cases} -\frac{1}{\sqrt{2}}, & t-1 \leq u < t; \\ \frac{1}{\sqrt{2}}, & t \leq u < t+1; \\ 0, & \text{otherwise} \end{cases}$$

Fig. 4: top row of plots

• to mine out info about other  $\lambda$ 's, form

$$\psi_{\lambda,t}^{\mathrm{H}}(u) \equiv \frac{1}{\sqrt{\lambda}} \psi^{\mathrm{H}}\left(\frac{u-t}{\lambda}\right) = \begin{cases} -\frac{1}{\sqrt{2\lambda}}, & t-\lambda \leq u < t; \\ \frac{1}{\sqrt{2\lambda}}, & t \leq u < t+\lambda; \\ 0, & \text{otherwise.} \end{cases}$$

Fig. 4: bottom row of plots

# Basics of Wavelet Analysis: V

- $\bullet$  can check that  $\psi^{\mathrm{H}}_{\lambda,t}(u)$  is a wavelet for all  $\lambda$  & t
- use  $\psi^{\mathrm{H}}_{\lambda,t}(u)$  to obtain

$$W^{\mathrm{H}}(\lambda,t) \equiv \int_{-\infty}^{\infty} \psi_{\lambda,t}^{\mathrm{H}}(u) x(u) \, du \propto D(\lambda,t)$$

left-hand side is Haar CWT

• can do the same with other wavelets:

$$W(\lambda,t) \equiv \int_{-\infty}^{\infty} \psi_{\lambda,t}(u) x(u) \, du$$
, where  $\psi_{\lambda,t}(u) \equiv \frac{1}{\sqrt{\lambda}} \psi\left(\frac{u-t}{\lambda}\right)$ 

left-hand side is CWT based on  $\psi(u)$ 

• interpretation for  $\psi^{\text{fdG}}(u)$  and  $\psi^{\text{Mh}}(u)$  (Fig. 1): differences of adjacent weighted averages

#### **Basics of Wavelet Analysis: VI**

• basic CWT result: if  $\psi(u)$  satisfies admissibility condition, can recover x(t) from its CWT:

$$x(t) = \frac{1}{C_{\psi}} \int_0^{\infty} \left[ \int_{-\infty}^{\infty} W(\lambda, t) \frac{1}{\sqrt{\lambda}} \psi\left(\frac{t-u}{\lambda}\right) \, du \right] \, \frac{d\lambda}{\lambda^2},$$

where  $C_{\psi}$  is constant depending just on  $\psi$ 

- conclusion:  $W(\lambda, t)$  equivalent to x(t)
- can also show that

$$\int_{-\infty}^{\infty} x^2(t) \, dt = \frac{1}{C_{\psi}} \left[ \int_0^{\infty} \int_{-\infty}^{\infty} W^2(\lambda, t) \, dt \right] \, \frac{d\lambda}{\lambda^2}$$

- LHS called energy in x(t)
- RHS integrand is energy density over  $\lambda \& t$
- Fig. 3: Mexican hat CWT of clock 571 data

## Beyond the CWT: the DWT

- critique: have transformed signal into an image
- can often get by with subsamples of  $W(\lambda, t)$
- leads to notion of discrete wavelet transform (DWT)
  - can regard as dyadic 'slices' through CWT
  - can further subsample slices at various t's
- DWT has appeal in its own right
  - most time series are sampled as discrete values (can be tricky to implement CWT)
  - can formulate as orthonormal transform (facilitates statistical analysis)
  - approximately decorrelates certain time series (including power law processes)
  - standardization to dyadic scales often adequate
  - can be faster than the fast Fourier transform!
- will concentrate on DWT for remainder of tutorial

## **Overview of DWT**

- let  $\mathbf{X} = [X_0, X_1, \dots, X_{N-1}]^T$  be observed time series (for convenience, assume N integer multiple of  $2^{J_0}$ )
- let  $\mathcal{W}$  be  $N \times N$  orthonormal DWT matrix
- $\mathbf{W} = \mathcal{W}\mathbf{X}$  is vector of DWT coefficients
- orthonormality says  $\mathbf{X} = \mathcal{W}^T \mathbf{W}$ , so  $\mathbf{X} \Leftrightarrow \mathbf{W}$
- can partition **W** as follows:

$$\mathbf{W} = egin{bmatrix} \mathbf{W}_1 \ dots \ \mathbf{W}_{J_0} \ \mathbf{V}_{J_0} \end{bmatrix}$$

- $\mathbf{W}_j$  contains  $N_j = N/2^j$  wavelet coefficients
  - related to changes of averages at scale  $\tau_j = 2^{j-1}$ ( $\tau_j$  is *j*th 'dyadic' scale)
  - related to times spaced  $2^j$  units apart
- $\mathbf{V}_{J_0}$  contains  $N_{J_0} = N/2^{J_0}$  scaling coefficients
  - related to averages at scale  $\lambda_{J_0} = 2^{J_0}$
  - related to times spaced  $2^{J_0}$  units apart

#### Example: Haar DWT

- Fig. 5:  $\mathcal{W}$  for Haar DWT with N = 16
  - first 8 rows yield  $\mathbf{W}_1 \propto changes$  on scale 1
  - next 4 rows yield  $\mathbf{W}_2 \propto changes$  on scale 2
  - next 2 rows yield  $\mathbf{W}_3 \propto changes$  on scale 4
  - next to last row yields  $\mathbf{W}_4 \propto change$  on scale 8
  - last row yields  $\mathbf{V}_4 \propto average$  on scale 16
- Fig. 6: Haar DWT coefficients for clock 571

#### **DWT** in Terms of Filters

• filter  $X_0, X_1, \ldots, X_{N-1}$  to obtain

$$2^{j/2}\widetilde{W}_{j,t} \equiv \sum_{l=0}^{L_j-1} h_{j,l} X_{t-l \bmod N}, \quad t = 0, 1, \dots, N-1$$

where  $h_{j,l}$  is *j*th level wavelet filter

- note: circular filtering

• subsample to obtain wavelet coefficients:

 $W_{j,t} = 2^{j/2} \widetilde{W}_{j,2^{j}(t+1)-1}, \quad t = 0, 1, \dots, N_{j} - 1,$ 

where  $W_{j,t}$  is the element of  $\mathbf{W}_j$ 

- Figs. 7 & 8: Haar, D(4), C(6) & LA(8) wavelet filters
- *j*th wavelet filter is band-pass with pass-band  $\left[\frac{1}{2^{j+1}}, \frac{1}{2^{j}}\right]$
- note: jth scale related to interval of frequencies
- similarly, scaling filters yield  $\mathbf{V}_{J_0}$
- Figs. 9 & 10: Haar, D(4), C(6) & LA(8) scaling filters
- $J_0$ th scaling filter is low-pass with pass-band  $[0, \frac{1}{2^{J_0+1}}]$

### Pyramid Algorithm: I

- can formulate DWT via 'pyramid algorithm'
  - elegant iterative algorithm for computing DWT
  - implicitly defines  ${\cal W}$
  - computes  $\mathbf{W} = \mathcal{W}\mathbf{X}$  using O(N) multiplications
    - \* 'brute force' method uses  $O(N^2)$
    - \* FFT algorithm uses  $O(N \log_2 N)$
- algorithm makes use of two basic filters
  - wavelet filter  $h_l$  of unit scale  $h_l \equiv h_{1,l}$
  - associated scaling filter  $g_l$

## The Wavelet Filter: I

- let  $h_l, l = 0, \ldots, L 1$ , be a real-valued filter
  - -L is filter width so  $h_0 \neq 0 \& h_{L-1} \neq 0$
  - -L must be even
  - assume  $h_l = 0$  for  $l < 0 \& l \ge L$
- $h_l$  called a wavelet filter if it has these 3 properties
  - 1. summation to zero:

$$\sum_{l=0}^{L-1} h_l = 0$$

2. unit energy:

$$\sum_{l=0}^{L-1} h_l^2 = 1$$

3. orthogonality to even shifts:

$$\sum_{l=0}^{L-1} h_l h_{l+2n} = \sum_{l=-\infty}^{\infty} h_l h_{l+2n} = 0$$

for all nonzero integers n

• 2 & 3 together called orthonormality property

### The Wavelet Filter: II

• transfer & squared gain functions for  $h_l$ :

$$H(f) \equiv \sum_{l=0}^{L-1} h_l e^{-i2\pi f l} \& \mathcal{H}(f) \equiv |H(f)|^2$$

• can argue that orthonormality property equivalent to

$$\mathcal{H}(f) + \mathcal{H}(f + \frac{1}{2}) = 2$$
 for all  $f$ 

- Fig. 11:  $\mathcal{H}(f)$  for Daubechies wavelet filters
  - -L = 2 case is Haar wavelet filter
  - filter cascade with averaging & differencing filters
  - high-pass filter with pass-band  $\left[\frac{1}{4}, \frac{1}{2}\right]$
  - can regard as half-band filter

### The Scaling Filter: I

- scaling filter:  $g_l \equiv (-1)^{l+1} h_{L-1-l}$ 
  - reverse  $h_l$  & flip sign of every other coefficient

- e.g.: 
$$h_0 = \frac{1}{\sqrt{2}} \& h_1 = -\frac{1}{\sqrt{2}} \Rightarrow g_0 = g_1 = \frac{1}{\sqrt{2}}$$
  
-  $g_l$  is 'quadrature mirror' filter for  $h_l$ 

- properties of  $h_l$  imply  $g_l$  has these properties:
  - 1. summation to  $\pm \sqrt{2}$ , so will assume

$$\sum_{l=0}^{L-1} g_l = \sqrt{2}$$

2. unit energy:

$$\sum_{l=0}^{L-1} g_l^2 = 1$$

3. orthogonality to even shifts:

$$\sum_{l=0}^{L-1} g_l g_{l+2n} = \sum_{l=-\infty}^{\infty} g_l g_{l+2n} = 0$$

for all nonzero integers n

4. orthogonality to wavelet filter at even shifts:

$$\sum_{l=0}^{L-1} g_l h_{l+2n} = \sum_{l=-\infty}^{\infty} g_l h_{l+2n} = 0$$

for all integers n

## The Scaling Filter: II

• transfer & squared gain functions for  $g_l$ :

$$G(f) \equiv \sum_{l=0}^{L-1} g_l e^{-i2\pi f l} \& \mathcal{G}(f) \equiv |G(f)|^2$$

• can argue that 
$$\mathcal{G}(f) = \mathcal{H}(f - \frac{1}{2})$$

- have  $\mathcal{G}(0) = \mathcal{H}(-\frac{1}{2}) = \mathcal{H}(\frac{1}{2}) \& \mathcal{G}(\frac{1}{2}) = \mathcal{H}(0)$
- since  $h_l$  is high-pass,  $g_l$  must be low-pass
- low-pass filter with pass-band  $[0, \frac{1}{4}]$
- can also regard as half-band filter
- orthonormality property equivalent to

$$\mathcal{G}(f) + \mathcal{G}(f + \frac{1}{2}) = 2 \text{ or } \mathcal{H}(f) + \mathcal{G}(f) = 2 \text{ for all } f$$

### Pyramid Algorithm: II

- define  $\mathbf{V}_0 \equiv \mathbf{X}$  and set j = 1
- input to *j*th stage of pyramid algorithm is  $\mathbf{V}_{j-1}$ 
  - $-\mathbf{V}_{j-1}$  is full-band
  - related to frequencies  $[0, \frac{1}{2^j}]$  in **X**
- filter with half-band filters and downsample:

$$W_{j,t} \equiv \sum_{l=0}^{L-1} h_l V_{j-1,2t+1-l \mod N_{j-1}}$$
$$V_{j,t} \equiv \sum_{l=0}^{L-1} g_l V_{j-1,2t+1-l \mod N_{j-1}},$$

 $t=0,\ldots,N_j-1$ 

- place these in vectors  $\mathbf{W}_j$  &  $\mathbf{V}_j$ 
  - $-\mathbf{W}_j$  are wavelet coefficients for scale  $\tau_j = 2^{j-1}$
  - $-\mathbf{V}_j$  are scaling coefficients for scale  $\lambda_j = 2^j$
- increment j and repeat above until  $j = J_0$
- yields DWT coefficients  $\mathbf{W}_1, \ldots, \mathbf{W}_{J_0}, \mathbf{V}_{J_0}$

### Pyramid Algorithm: III

- can formulate inverse pyramid algorithm (recovers  $\mathbf{V}_{j-1}$  from  $\mathbf{W}_j$  and  $\mathbf{V}_j$ )
- $\bullet$  algorithm implicitly defines transform matrix  ${\cal W}$
- partition  $\mathcal{W}$  commensurate with  $\mathbf{W}_j$ :

$$\mathcal{W} = \begin{bmatrix} \mathcal{W}_1 \\ \mathcal{W}_2 \\ \vdots \\ \mathcal{W}_{J_0} \\ \mathcal{V}_{J_0} \end{bmatrix} \text{ parallels } \mathbf{W} = \begin{bmatrix} \mathbf{W}_1 \\ \mathbf{W}_2 \\ \vdots \\ \mathbf{W}_{J_0} \\ \mathbf{V}_{J_0} \end{bmatrix}$$

• rows of  $\mathcal{W}_j$  use *j*th level filter  $h_{j,l}$  with DFT

$$H(2^{j-1}f)\prod_{l=0}^{j-2}G(2^lf)$$

 $(h_{j,l} \text{ has } L_j = (2^j - 1)(L - 1) + 1 \text{ nonzero elements})$ 

•  $\mathcal{W}_j$  is  $N_j \times N$  matrix such that

 $\mathbf{W}_j = \mathcal{W}_j \mathbf{X}$  and  $\mathcal{W}_j \mathcal{W}_j^T = I_{N_j}$ 

# **Two Consequences of Orthonormality**

• multiresolution analysis (MRA)

$$\mathbf{X} = \mathcal{W}^T \mathbf{W} = \sum_{j=1}^{J_0} \mathcal{W}_j^T \mathbf{W}_j + \mathcal{V}_{J_0}^T \mathbf{V}_{J_0} \equiv \sum_{j=1}^{J_0} \mathcal{D}_j + \mathcal{S}_{J_0}$$

- scale-based additive decomposition
- $-\mathcal{D}_{j}$ 's &  $\mathcal{S}_{J_{0}}$  called details & smooth
- analysis of variance
  - consider 'energy' in time series:

$$\|\mathbf{X}\|^2 = \mathbf{X}^T \mathbf{X} = \sum_{t=0}^{N-1} X_t^2$$

- energy preserved in DWT coefficients:

$$\|\mathbf{W}\|^2 = \|\mathcal{W}\mathbf{X}\|^2 = \mathbf{X}^T \mathcal{W}^T \mathcal{W}\mathbf{X} = \mathbf{X}^T \mathbf{X} = \|\mathbf{X}\|^2$$

- since  $\mathbf{W}_1, \ldots, \mathbf{W}_{J_0}, \mathbf{V}_{J_0}$  partitions  $\mathbf{W}$ , have

$$\|\mathbf{W}\|^2 = \sum_{j=1}^{J_0} \|\mathbf{W}_j\|^2 + \|\mathbf{V}_{J_0}\|^2,$$

leading to analysis of sample variance:

$$\hat{\sigma}^2 \equiv \frac{1}{N} \sum_{t=0}^{N-1} \left( X_t - \overline{X} \right)^2 = \frac{1}{N} \sum_{j=1}^{J_0} \|\mathbf{W}_j\|^2 + \left( \frac{1}{N} \|\mathbf{V}_{J_0}\|^2 - \overline{X}^2 \right)$$

- scale-based decomposition (cf. frequency-based)

#### Variation: Maximal Overlap DWT

• can eliminate downsampling and use

$$\widetilde{W}_{j,t} \equiv \frac{1}{2^{j/2}} \sum_{l=0}^{L_j - 1} h_{j,l} X_{t-l \bmod N}, \quad t = 0, 1, \dots, N - 1$$

to define MODWT coefficients  $\widetilde{\mathbf{W}}_j$  (& also  $\widetilde{\mathbf{V}}_j$ )

- unlike DWT, MODWT is not orthonormal (in fact MODWT is highly redundant)
- like DWT, can do MRA & analysis of variance:

$$\|\mathbf{X}\|^{2} = \sum_{j=1}^{J_{0}} \|\widetilde{\mathbf{W}}_{j}\|^{2} + \|\widetilde{\mathbf{V}}_{J_{0}}\|^{2}$$

- unlike DWT, MODWT works for all samples sizes N (i.e., power of 2 assumption is not required)
  - if N is power of 2, can compute MODWT using  $O(N \log_2 N)$  operations (i.e., same as FFT algorithm)
  - contrast to DWT, which uses O(N) operations
- Fig. 12: Haar MODWT coefficients for clock 571 (cf. Fig. 6 with DWT coefficients)

#### **Definition of Wavelet Variance**

- let  $X_t, t = \ldots, -1, 0, 1, \ldots$ , be a stochastic process
- run  $X_t$  through *j*th level wavelet filter:

$$\overline{W}_{j,t} \equiv \sum_{l=0}^{L_j-1} \tilde{h}_{j,l} X_{t-l}, \quad t = \dots, -1, 0, 1, \dots,$$

which should be contrasted with

$$\widetilde{W}_{j,t} \equiv \sum_{l=0}^{L_j-1} \widetilde{h}_{j,l} X_{t-l \bmod N}, \quad t = 0, 1, \dots, N-1$$

• definition of time dependent wavelet variance (also called wavelet spectrum):

$$\nu_{X,t}^2(\tau_j) \equiv \operatorname{var} \{ \overline{W}_{j,t} \},\,$$

assuming var  $\{\overline{W}_{j,t}\}$  exists and is finite

- $\nu_{X,t}^2(\tau_j)$  depends on  $\tau_j$  and t
- will consider time independent wavelet variance:

$$\nu_X^2(\tau_j) \equiv \operatorname{var}\left\{\overline{W}_{j,t}\right\}$$

(can be easily adapted to time varying situation)

# **Rationale for Wavelet Variance**

- decomposes variance on scale by scale basis
- useful substitute/complement for spectrum
- useful substitute for process/sample variance

#### Variance Decomposition

• suppose  $X_t$  has power spectrum  $S_X(f)$ :

$$\int_{-1/2}^{1/2} S_X(f) \, df = \operatorname{var} \{ X_t \};$$

i.e., decomposes var  $\{X_t\}$  across frequencies f

- involves uncountably infinite number of f's
- $-S_X(f)\Delta f \approx \text{contribution to var} \{X_t\} \text{ due to } f \text{'s}$ in interval of length  $\Delta f$  centered at f
- wavelet variance analog to fundamental result:

$$\sum_{j=1}^{\infty} \nu_X^2(\tau_j) = \operatorname{var} \{X_t\}$$

i.e., decomposes var  $\{X_t\}$  across scales  $\tau_j$ 

- recall DWT/MODWT and sample variance
- involves countably infinite number of  $\tau_j$ 's
- $-\nu_X^2(\tau_j)$  contribution to var  $\{X_t\}$  due to scale  $\tau_j$
- $-\nu_X(\tau_j)$  has same units as  $X_t$  (easier to interpret)

#### Spectrum Substitute/Complement

• because  $\tilde{h}_{j,l} \approx$  bandpass over  $[1/2^{j+1}, 1/2^j]$ ,

$$\nu_X^2(\tau_j) \approx 2 \int_{1/2^{j+1}}^{1/2^j} S_X(f) \, df$$

- if  $S_X(f)$  'featureless', info in  $\nu_X^2(\tau_j) \Leftrightarrow \inf S_X(f)$
- $\nu_X^2(\tau_j)$  more succinct: only 1 value per octave band
- example:  $S_X(f) \propto |f|^{\alpha}$ , i.e., power law process
  - can deduce  $\alpha$  from slope of log  $S_X(f)$  vs. log f
  - implies  $\nu_X^2(\tau_j) \propto \tau_j^{-\alpha-1}$  approximately
  - can deduce  $\alpha$  from slope of log  $\nu_X^2(\tau_j)$  vs. log  $\tau_j$
  - no loss of 'info' using  $\nu_X^2(\tau_j)$  rather than  $S_X(f)$
- with Haar wavelet, obtain pilot spectrum estimate proposed in Blackman & Tukey (1958)

# Substitute for Variance: I

- can be difficult to estimate process variance
- $\nu_X^2(\tau_j)$  useful substitute: easy to estimate & finite

• let 
$$\mu = E\{X_t\}$$
 be known,  $\sigma^2 = \operatorname{var}\{X_t\}$  unknown

• can estimate  $\sigma^2$  using

$$\tilde{\sigma}^2 \equiv \frac{1}{N} \sum_{t=0}^{N-1} (X_t - \mu)^2$$

• estimator above is unbiased:  $E\{\tilde{\sigma}^2\} = \sigma^2$ 

• if  $\mu$  is unknown, can estimate  $\sigma^2$  using

$$\hat{\sigma}^2 \equiv \frac{1}{N} \sum_{t=0}^{N-1} (X_t - \overline{X})^2$$

• there is some (non-pathological!)  $X_t$  such that

$$\frac{E\{\hat{\sigma}^2\}}{\sigma^2} < \epsilon$$

for any gvien  $\epsilon > 0$  &  $N \ge 1$ 

- $\hat{\sigma}^2$  can badly underestimate  $\sigma^2$ !
- example: power law process with  $-1 < \alpha < 0$

#### Substitute for Variance: II

- Q: why is wavelet variance useful when  $\sigma^2$  is not?
- replaces 'global' variability with variability over scales
- if  $X_t$  stationary with mean  $\mu$ , then

$$E\{\overline{W}_{j,t}\} = \sum_{l=0}^{L_j-1} \tilde{h}_{j,l} E\{X_{t-l}\} = \mu \sum_{l=0}^{L_j-1} \tilde{h}_{j,l} = 0$$

because  $\Sigma_l \tilde{h}_{j,l} = 0$ 

- $E{\overline{W}_{j,t}}$  known, so can get unbiased estimator of var  ${\overline{W}_{j,t}} = \nu_X^2(\tau_j)$
- certain nonstationary  $X_t$  have well-defined  $\nu_X^2(\tau_j)$
- example: power law processes with  $\alpha \leq -1$  (example of process with stationary increments)

#### Estimation of Wavelet Variance: I

• can base estimator on MODWT of  $X_0, X_1, \ldots, X_{N-1}$ :

$$\widetilde{W}_{j,t} \equiv \sum_{l=0}^{L_j-1} \widetilde{h}_{j,l} X_{t-l \mod N}, \quad t = 0, 1, \dots, N-1$$

(DWT-based estimator possible, but less efficient)

• recall that

$$\overline{W}_{j,t} \equiv \sum_{l=0}^{L_j-1} \tilde{h}_{j,l} X_{t-l}, \quad t = 0, \pm 1, \pm 2, \dots$$

so  $\widetilde{W}_{j,t} = \overline{W}_{j,t}$  if mod not needed:  $L_j - 1 \le t < N$ 

• if  $N - L_j \ge 0$ , unbiased estimator of  $\nu_X^2(\tau_j)$  is

$$\hat{\nu}_X^2(\tau_j) \equiv \frac{1}{N - L_j + 1} \sum_{t=L_j - 1}^{N-1} \widetilde{W}_{j,t}^2 = \frac{1}{M_j} \sum_{t=L_j - 1}^{N-1} \overline{W}_{j,t}^2,$$

where  $M_j \equiv N - L_j + 1$ 

• can also construct biased estimator of  $\nu_X^2(\tau_j)$ :

$$\tilde{\nu}_X^2(\tau_j) \equiv \frac{1}{N} \sum_{t=0}^{N-1} \widetilde{W}_{j,t}^2 = \frac{1}{N} \Big( \sum_{t=0}^{L_j-2} \widetilde{W}_{j,t}^2 + \sum_{t=L_j-1}^{N-1} \overline{W}_{j,t}^2 \Big)$$

1st sum in parentheses influenced by circularity

# Estimation of Wavelet Variance: II

- biased estimator unbiased if  $\{X_t\}$  white noise
- biased estimator offers exact analysis of  $\hat{\sigma}^2$ ; unbiased estimator need not
- biased estimator can have better mean square error (Greenhall *et al.*, 1999; need to 'reflect'  $X_t$ )

# Statistical Properties of $\hat{\nu}_X^2(\tau_j)$

- suppose  $\{\overline{W}_{j,t}\}$  Gaussian, mean 0 & spectrum  $S_j(f)$
- suppose square integrability condition holds:

$$A_j \equiv \int_{-1/2}^{1/2} S_j^2(f) \, df < \infty \& S_j(f) > 0$$

(holds for power law processes if L large enough)

- can show  $\hat{\nu}_X^2(\tau_j)$  asymptotically normal with mean  $\nu_X^2(\tau_j)$  & large sample variance  $2A_j/M_j$
- can estimate  $A_j$  and use with  $\hat{\nu}_X^2(\tau_j)$ to construct confidence interval for  $\nu_X^2(\tau_j)$
- example
  - Fig. 13: clock errors  $X_t \equiv X_t^{(0)}$  along with differences  $X_t^{(i)} \equiv X_t^{(i-1)} X_{t-1}^{(i-1)}$  for i = 1, 2
  - Fig. 14:  $\hat{\nu}_X^2(\tau_j)$  for clock errors
  - Fig. 15:  $\hat{\nu}_{\overline{Y}}^2(\tau_j)$  for  $\overline{Y}_t \propto X_t^{(1)}$
  - Haar  $\hat{\nu}_{\overline{Y}}^2(\tau_j)$  related to Allan variance  $\sigma_{\overline{Y}}^2(2,\tau_j)$ :

$$\nu_{\overline{Y}}^2(\tau_j) = \frac{1}{2}\sigma_{\overline{Y}}^2(2,\tau_j)$$

#### **Decorrelation of FD Processes**

•  $X_t$  'fractionally differenced' if its spectrum is

$$S_X(f) = \frac{\sigma_\epsilon^2}{|2\sin(\pi f)|^{2\delta}},$$

where  $\sigma_{\epsilon}^2 > 0$  and  $-\frac{1}{2} < \delta < \frac{1}{2}$ 

- note: for small f, have  $S_X(f) \approx C/|f|^{2\delta}$ ; i.e., power law with  $\alpha = -2\delta$
- if  $\delta = 0$ , FD process is white noise
- if  $0 < \delta < \frac{1}{2}$ , FD stationary with 'long memory'
- can extend definition to  $\delta \geq \frac{1}{2}$ 
  - nonstationary 1/f type process
  - also called ARFIMA( $0,\delta,0$ ) process
- Fig. 16: DWT of simulated FD process,  $\delta = 0.4$  (sample autocorrelation sequences (ACSs) on right)

### **DWT** as Whitening Transform

- sample ACSs suggest  $\mathbf{W}_j \approx$  uncorrelated
- since FD process is stationary, so are  $\mathbf{W}_j$  (ignoring terms influenced by circularity)
- Fig. 17: spectra for  $\mathbf{W}_j$ , j = 1, 2, 3, 4
- $\mathbf{W}_{j}$  &  $\mathbf{W}_{j'}, j \neq j'$ , approximately uncorrelated (approximation improves as L increases)
- DWT thus acts as a whitening transform
- lots of uses for whitening property, including:
  - 1. testing for variance changes
  - 2. bootstrapping time series statistics
  - 3. estimating  $\delta$  for stationary/nonstationary fractional difference processes with trend

#### Estimation for FD Processes: I

- extension of work by Wornell; McCoy & Walden
- problem: estimate  $\delta$  from time series  $U_t$  such that

$$U_t = T_t + X_t$$

where

- $-T_t \equiv \sum_{j=0}^r a_j t^j$  is polynomial trend
- $-X_t$  is FD process, but can have  $\delta \geq \frac{1}{2}$
- DWT wavelet filter of width L has embedded differencing operation of order L/2
- if  $\frac{L}{2} \ge r+1$ , reduces polynomial trend to 0
- can partition DWT coefficients as

$$\mathbf{W} = \mathbf{W}_s + \mathbf{W}_b + \mathbf{W}_w$$

where

- $-\mathbf{W}_s$  has scaling coefficients and 0s elsewhere
- $-\mathbf{W}_s$  has boundary-dependent wavelet coefficients
- $-\mathbf{W}_w$  has boundary-independent wavelet coefficients

# Estimation for FD Processes: II

• since  $\mathbf{U} = \mathcal{W}^T \mathbf{W}$ , can write

$$\mathbf{U} = \mathcal{W}^T(\mathbf{W}_s + \mathbf{W}_b) + \mathcal{W}^T\mathbf{W}_w \equiv \widehat{\mathbf{T}} + \widehat{\mathbf{X}}$$

- Fig. 18: example with fractional frequency deviates
- can use values in  $\mathbf{W}_w$  to form likelihood:

$$L(\delta, \sigma_{\epsilon}^{2}) \equiv \prod_{j=1}^{J_{0}} \prod_{t=1}^{N'_{j}} \frac{1}{\left(2\pi\sigma_{j}^{2}\right)^{1/2}} e^{-W_{j,t+L'_{j}-1}^{2}/(2\sigma_{j}^{2})}$$

where

$$\sigma_j^2 \equiv \int_{-1/2}^{1/2} \mathcal{H}_j(f) \frac{\sigma_\epsilon^2}{|2\sin(\pi f)|^{2\delta}} df;$$

and  $\mathcal{H}_j(f)$  is squared gain for  $h_{j,l}$ 

- leads to maximum likelihood estimator  $\hat{\delta}$  for  $\delta$
- works well in Monte Carlo simulations
- get  $\hat{\delta} \doteq 0.39 \pm 0.03$  for fractional frequency deviates

#### **DWT-based Signal Extraction: I**

- DWT analysis of  $\mathbf{X}$  yields  $\mathbf{W} = \mathcal{W}\mathbf{X}$
- DWT synthesis  $\mathbf{X} = \mathcal{W}^T \mathbf{W}$  yields
  - multiresolution analysis (MRA)
  - -estimator of 'signal'  ${\bf D}$  hidden in  ${\bf X}:$ 
    - \* modify **W** to get **W**'
    - $\ast$  use  $\mathbf{W}'$  to form signal estimate:

$$\widehat{\mathbf{D}} \equiv \mathcal{W}^T \mathbf{W}'$$

- key ideas behind wavelet-based signal estimation
  - DWT can isolate signals in small number of  $W_n$ 's
  - can 'threshold' or 'shrink'  $W_n$ 's
- key ideas lead to 'waveshrink' (Donoho and Johnstone, 1995)

#### **DWT-based Signal Extraction: II**

- thresholding schemes involve
  - 1. computing  $\mathbf{W} \equiv \mathcal{W} \mathbf{X}$
  - 2. defining  $\mathbf{W}^{(t)}$  as vector with *n*th element  $W_n^{(t)} = \begin{cases} 0, & \text{if } |W_n| \le \delta; \\ \text{some nonzero value, otherwise,} \end{cases}$

where nonzero values are yet to be defined

3. estimating  $\mathbf{D}$  via  $\widehat{\mathbf{D}}^{(t)} \equiv \mathcal{W}^T \mathbf{W}^{(t)}$ 

• simplest scheme is 'hard thresholding:'

$$W_n^{(ht)} = \begin{cases} 0, & \text{if } |W_n| \le \delta; \\ W_n, & \text{otherwise.} \end{cases}$$

Fig. 19: solid line ('kill/keep' strategy)

• alterative scheme is 'soft thresholding:'

$$W_n^{(st)} = \text{sign} \{W_n\} (|W_n| - \delta)_+,$$

where

sign 
$$\{W_n\} \equiv \begin{cases} +1, & \text{if } W_n > 0; \\ 0, & \text{if } W_n = 0; \\ -1, & \text{if } W_n < 0. \end{cases}$$
 and  $(x)_+ \equiv \begin{cases} x, & \text{if } x \ge 0; \\ 0, & \text{if } x < 0. \end{cases}$ 

Fig. 19: dashed line

## **DWT-based Signal Extraction: III**

• third scheme is 'mid thresholding:'

$$W_n^{(mt)} = \operatorname{sign} \{W_n\} (|W_n| - \delta)_{++},$$

where

$$(|W_n| - \delta)_{++} \equiv \begin{cases} 2(|W_n| - \delta)_+, & \text{if } |W_n| < 2\delta; \\ |W_n|, & \text{otherwise} \end{cases}$$

Fig. 19: dotted line

- Q: how should  $\delta$  be set?
- A: universal' threshold (Donoho & Johnstone, 1995) (lots of other answers have been proposed)
  - specialize to model  $\mathbf{X} = \mathbf{D} + \boldsymbol{\epsilon}$ , where  $\boldsymbol{\epsilon}$  is Gaussian white noise with variance  $\sigma_{\epsilon}^2$
  - 'universal' threshold:  $\delta_U \equiv \sqrt{[2\sigma_\epsilon^2 \log(N)]}$
  - rationale for  $\delta_{\mathbf{U}}$ :
    - \* suppose  $\mathbf{D} = \mathbf{0}$  & hence  $\mathbf{W}$  is white noise also \* as  $N \to \infty$ , have

$$\mathbf{P}\left[\max_{n}|W_{n}|\leq\delta_{\mathbf{U}}\right]\rightarrow1$$

so all  $\mathbf{W}^{(ht)} = 0$  with high probability

\* will estimate correct **D** with high probability

# **DWT-based Signal Extraction: IV**

• can estimate  $\sigma_{\epsilon}^2$  using median absolute deviation (MAD):

$$\hat{\sigma}_{(\text{MAD})} \equiv \frac{\text{median}\left\{|W_{1,0}|, |W_{1,1}|, \dots, |W_{1,\frac{N}{2}-1}|\right\}}{0.6745},$$

where  $W_{1,t}$ 's are elements of  $\mathbf{W}_1$ 

- Fig. 20: application to NMR series
- has potential application in dejamming GPS signals (with roles of 'signal' and 'noise' swapped!)

### Web Material and Books

• Wavelet Digest

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http://www.wavelet.org/
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• MathSoft's wavelet resource page

http://www.mathsoft.com/wavelets.html

- books
  - R. Carmona, W.–L. Hwang & B. Torrésani (1998), *Practical Time-Frequency Analysis*, Academic Press
  - S. G. Mallat (1999), A Wavelet Tour of Signal Processing (Second Edition), Academic Press
  - R. T. Ogden (1997), Essential Wavelets for Statistical Applications and Data Analysis, Birkhäuser
  - D. B. Percival & A. T. Walden (2000), Wavelet Methods for Time Series Analysis, Cambridge University Press (will appear in July/August) http://www.staff.washington.edu/dbp/wmtsa.html
  - B. Vidakovic (1999), Statistical Modeling by Wavelets, John Wiley & Sons.

### Software

• Matlab

- WAVELAB (free):

http://www-stat.stanford.edu/~wavelab

- WAVEBOX (commercial):

http://www.toolsmiths.com/

• Mathcad Wavelets Extension Pack (commercial):

http://www.mathsoft.com/mathcad/ebooks/wavelets.asp

- S-Plus software
  - WAVETHRESH (free):

http://lib.stat.cmu.edu/S/wavethresh

- S+WAVELETS (commercial):

http://www.mathsoft.com/splsprod/wavelets.html